

Math 5654 Midterm 2 Solutions

Problem 1

Let $Z = \Pi_{\mathcal{L}}X$.

We have $X = cY$, where $X - cY \perp Y$. Then $\langle X, Y \rangle - c\langle Y, Y \rangle = 0$, so $3 - c4 = 0$, or $c = 3/4$.

Problem 2

Let $Y = (Y_1, Y_2)$.

(i)

$$\Sigma_{Y,Y} = \begin{bmatrix} \text{Cov}(Y_1, Y_1) & \text{Cov}(Y_1, Y_2) \\ \text{Cov}(Y_2, Y_1) & \text{Cov}(Y_2, Y_2) \end{bmatrix} = \begin{bmatrix} 2 & 1 \\ 1 & 1 \end{bmatrix}.$$

(ii) By Problem 4.3.8 in the text,

$$a_1Y_1 + a_2Y_2 = AY,$$

where

$$\begin{aligned} A &= Q_{XY}Q_{YY}^{-1} = \begin{bmatrix} 3 & 1 \end{bmatrix} \begin{bmatrix} 2 & 1 \\ 1 & 1 \end{bmatrix}^{-1} \\ &= \begin{bmatrix} 3 & 1 \end{bmatrix} \begin{bmatrix} 1 & -1 \\ -1 & 2 \end{bmatrix} = \begin{bmatrix} 2 & -1 \end{bmatrix}. \end{aligned}$$

Thus $a_1 = 2$ and $a_2 = -1$.

More direct approach We can solve two equations: $X - a_1Y_1 - a_2Y_2 \perp Y_1$, $X - a_1Y_1 - a_2Y_2 \perp Y_2$, which gives us

$$\begin{cases} 3 - 2a_1 - a_2 = 0, \\ 1 - a_1 - a_2 = 0. \end{cases}$$

Problem 3

By Definition 4.2.18, every linear combination of X_1, \dots, X_m is a one-dimensional Gaussian random variable.

Since each element Y_j is a linear combination of X_1, \dots, X_m , every linear combination of Y_1, \dots, Y_n is a linear combination of X_1, \dots, X_m . Thus every linear combination of Y_1, \dots, Y_n is a one-dimensional Gaussian random variable.

By Definition 4.2.18, (Y_1, \dots, Y_n) is Gaussian.

Problem 4

(i) Since $\mathcal{L}_1 \subset \mathcal{L}_2$, $Z \in \mathcal{L}_2$. So Z cannot be a better approximation than W , the best approximation in \mathcal{L}_2 . That is, $\|X - W\| \leq \|X - Z\|$.

(ii) When (X, Y_1, Y_2) is Gaussian, we know by Theorem 4.2.25 that

$$E(X|Y_1, Y_2) = Z.$$

Thus Z is the best approximation to X in all of $\sigma(Y_1, Y_2)$. Since $\mathcal{L}_2 \subset \sigma(Y_1, Y_2)$, Z is the best approximation to X in \mathcal{L}_2 . That is, $Z = W$.

Problem 5

Consider a nonnegative function f on \mathbb{R}^{n+2} .

Then

$$\begin{aligned} & Ef(X_0, \dots, X_{n+1}) \\ &= \int f(x_0, \dots, x_{n+1}) \pi_0(x_0) p(x_0, x_1) \dots p(x_{n-1}, x_n) p(x_n, x_{n+1}) dx_0 \dots dx_{n+1} \\ &= \int_{\mathbb{R}^{n+1}} \left(\int_{\mathbb{R}} f(x_0, \dots, x_n, x_{n+1}) p(x_n, x_{n+1}) dx_{n+1} \right) \\ &\quad \pi_0(x_0) p(x_0, x_1) \dots p(x_{n-1}, x_n) dx_0 \dots dx_n \\ &= E \left(\int_{\mathbb{R}} f(X_0, \dots, X_n, x_{n+1}) p(X_n, x_{n+1}) dx_{n+1} \right). \end{aligned}$$

By Definition 5.1.2, (X_n) is a time-homogeneous Markov chain with transition kernel p .

Problem 6

(i) When h is larger, the *relative* size of the noise is smaller.

(ii) In the notation of Problem 5.2.3, $v^2 = h^2 + 1$.

By Problem 5.2.3, the limit of $\bar{\sigma}_n^2$ is given by x , where x is the unique *positive* solution of

$$x = \frac{1}{h^2 + 1} \frac{x}{h^2 x + h^2 + 1} + \frac{1}{h^2 + 1}.$$

Simplifying, we have

$$(h^2 + 1)x = \frac{x}{h^2 x + h^2 + 1} + 1.$$

This gives

$$h^2 (h^2 + 1)x^2 + (h^2 + 1)^2 x = x + h^2 x + h^2 + 1.$$

Dividing both sides by $h^2 + 1$ we have

$$h^2 x^2 + (h^2 + 1)x = x + 1,$$

or

$$h^2 x^2 + h^2 x - 1 = 0,$$

i.e.

$$x^2 + x - \frac{1}{h^2} = 0.$$

Hence

$$x = \frac{-1 + \sqrt{1 + \frac{4}{h^4}}}{2}.$$

This quantity decreases as h increases, which is consistent with (i).

Problem 7

First, a remark. For any nonnegative function h on \mathbb{R}^{n+1} ,

$$\begin{aligned} Eh(X_0, \dots, X_n) &= Eh(X_0, X_0U_1, \dots, X_0U_1 \dots U_n) \\ &= \int_{\mathbb{R} \times [0,1]^n} f(x_0, x_0u_1, \dots, x_0u_1 \dots u_n) \pi_0(x_0) dx_0 du_1 \dots du_n. \end{aligned}$$

Now consider a nonnegative function f on \mathbb{R}^{n+2} .

Then

$$\begin{aligned} Ef(X_0, \dots, X_{n+1}) &= Ef(X_0, X_0U_1, \dots, X_0U_1 \dots U_{n+1}) \\ &= \int_{\mathbb{R} \times [0,1]^{n+1}} f(x_0, x_0u_1, \dots, x_0u_1 \dots u_{n+1}) \pi_0(x_0) dx_0 du_1 \dots du_{n+1} \\ &= \int_{\mathbb{R} \times [0,1]^n} \left(\int_{[0,1]} f(x_0, x_0u_1, \dots, x_0u_1 \dots u_{n+1}) du_{n+1} \right) \pi_0(x_0) dx_0 du_1 \dots du_n. \end{aligned}$$

Now we apply the remark, with

$$h(x_0, \dots, x_n) = \int_{[0,1]} f(x_0, x_0u_1, \dots, x_0u_1 \dots u_{n+1}) du_{n+1},$$

and we see that

$$Ef(X_0, \dots, X_{n+1}) = E \left(\int_{[0,1]} f(X_0, X_1, \dots, X_n, X_n u_{n+1}) du_{n+1} \right).$$

Making the substitution $x_n u_{n+1} = x_{n+1}$, $X_n du_{n+1} = dx_{n+1}$, this becomes

$$\begin{aligned} Ef(X_0, \dots, X_{n+1}) &= E \left(\int_{[0, X_n]} f(X_0, X_1, \dots, X_n, x_{n+1}) \frac{1}{X_n} dx_{n+1} \right) \\ &= E \left(\int_{\mathbb{R}} f(X_0, X_1, \dots, X_n, x_{n+1}) \frac{1}{X_n} I_{[0, X_n]}(x_{n+1}) dx_{n+1} \right). \end{aligned}$$

By Definition 5.1.2, (X_n) is a time-homogeneous Markov chain with transition kernel p , where

$$p(x, y) = \frac{1}{x} I_{[0, x]}(y).$$

Alternate approach used by one student Let $r(y) = I_{[0,1]}(y)$, so that r is the density of U_n for each n .

Let $\rho_n(x_0, \dots, x_n)$ be the density of (X_0, \dots, X_n) , assuming that it exists. Then the joint density of $(X_0, \dots, X_n, U_{n+1})$ is $\rho_n(x_0, \dots, x_n) r(u_{n+1})$.

For any nonnegative function f on \mathbb{R}^{n+2} ,

$$\begin{aligned} Ef(X_0, X_1, \dots, X_{n+1}) &= Ef(X_0, X_1, \dots, X_n, X_n U_{n+1}) \\ &= \int f(x_0, \dots, x_n, x_n u_{n+1}) \rho(x_0, \dots, x_n) r(u_{n+1}) dx_0 \dots, dx_n du_{n+1} \\ &= \int f(x_0, \dots, x_n, x_n u_{n+1}) \rho(x_0, \dots, x_n) r\left(\frac{x_{n+1}}{x_n}\right) dx_0 \dots, dx_n \left(\frac{1}{x_n}\right) dx_{n+1} \\ &= \int f(x_0, \dots, x_n, x_n u_{n+1}) \rho(x_0, \dots, x_n) p(x_n, x_{n+1}) dx_0 \dots, dx_n dx_{n+1}, \end{aligned}$$

where $p(x_n, x_{n+1}) = \left(\frac{1}{x_n}\right) r\left(\frac{x_{n+1}}{x_n}\right)$. Hence

$$\rho_{n+1}(x_0, \dots, x_{n+1}) = \rho_n(x_0, \dots, x_n) p(x_n, x_{n+1}).$$

Considering $n = 0, 1, \dots$, it follows that $\rho_n(x_0, \dots, x_n)$ exists for all n and is given by

$$\rho_n(x_0, \dots, x_n) = \pi_0(x_0) p(x_0, x_1) \dots p(x_{n-1}, x_n).$$

Thus Problem 5 of the exam is applicable.