

# Solutions: Math8651 - #7 (Fall 2002)

1. Fix a constant  $A$ , and set  $X_i^A = X_i 1_{|X_i| \leq A}$ . Note that  $1/n E(\max_{i < n} |X_i^A|) \leq A$ . On the other hand,

$$E(\max_{i \leq n} |X_i|/n) \leq n^{-1} E(\max_{i \leq n} |X_i^A|) + n^{-1} E \sum_{i=1}^n E(|X_i - X_i^A|).$$

Note that  $E(|X_i - X_i^A|) = E(|X_1 - X_1^A|) \rightarrow_{i \rightarrow \infty} 0$  by dominated convergence, and hence

$$\limsup E(\max_{i \leq n} |X_i|/n) \leq A.$$

$A$  being arbitrary, the claim follows.

2. First, note that as a consequence of Fubini's theorem, and the fact that the probability distribution function is continuous,

$$P(X_i = X_j) = \int P(X_j = z) dP(z) = 0. \quad (1)$$

Next, using the exchangeability of the  $X_i$ 's,

$$P(A_n) = P(X_n > \max_{k \neq n} X_k) = P(X_i > \max_{k \neq i} X_k) =: P(B_{i,n})$$

while, by (1),  $1 = \sum_{i=1}^n P(B_{i,n})$ . Hence,  $nP(A_n) = 1$ .

Also, for any indices  $i_1 < i_2 < \dots < i_k$ , and  $i_0 = 1$ ,

$$P(\cap_{j=1}^k A_{i_j}) = P(\cap_{j=1}^k \{X_{i_j} > \max_{i_{j-1} \leq \ell < i_j} X_\ell\}) := P(\mathbf{X}_i \in B_k)$$

where  $\mathbf{i} = (i_1, \dots, i_k)$  and for any multi index  $\mathbf{j} = (j_1, \dots, j_{i_k})$ ,  $\mathbf{X}_j = (X_{j_1}, X_{j_2}, \dots, X_{j_{i_k}})$ . Let  $\pi$  be a permutation of the indices  $1, \dots, i_k$ . Then, by exchangeability,  $P(\mathbf{X}_{\pi \mathbf{i}} \in B_k) = P(\mathbf{X}_i \in B_k)$ .

Pick now a permutation at random, uniformly over the  $i_k!$  permutations, and apply Fubini's theorem:

$$\frac{1}{i_k!} \sum_{\pi} P(\mathbf{X}_{\pi \mathbf{i}} \in B_k) = E\left(\frac{1}{i_k!} \sum_{\pi} 1_{\mathbf{X}_{\pi \mathbf{i}} \in B_k}\right). \quad (2)$$

So we need to count, for a given  $(X_1, \dots, X_{i_k})$  with no two entries equal, the number  $N_k$  of permutations that make

$$\mathbf{X}_{\pi \mathbf{i}} \in B_k. \quad (3)$$

Fix as  $\pi_0$  the permutation such that  $X_{\pi_0 1} < X_{\pi_0 2} < \dots < X_{\pi_0 i_k}$ . In order for a permutation  $\pi$  to satisfy (3), one has to have  $\pi(i_k) = \pi_0(i_k)$ , and one has to have that  $(X_{\pi 1}, \dots, X_{\pi i_{k-1}}) \in B_{k-1}$ . There are  $\binom{i_k - i_{k-1} - 1}{i_k - 1}$

different ways to split the  $i_k - 1$  indices into groups of size  $i_k - i_{k-1} - 1$  and  $i_{k-1}$ . So the number of permutations needed is

$$N_k = (i_k - i_{k-1} - 1)! \binom{i_k - i_{k-1} - 1}{i_k - 1} N_{k-1} = \frac{(i_k - 1)!}{i_{k-1}!} N_{k-1}$$

Iterating, one gets

$$N_k = (i_k - 1)! \prod_{j=1}^{k-1} \frac{(i_j - 1)!}{i_j!} = \frac{(i_k - 1)!}{\prod_{j=1}^{k-1} i_j}$$

Substituting in (2), one gets

$$P(\mathbf{X}_i \in B_k) = \frac{1}{\prod_{j=1}^k i_j} = \prod_{j=1}^k P(A_{i_j}).$$

*Claim:*  $P(A_i \text{ infinitely often}) = 1$  This follows from the independence and the Borel-Cantelli lemma, for  $\sum 1/i = \infty$ .

Finally,  $T = \sum_{j=2}^{\infty} j 1_{A_j \cap_{i=2}^{j-1} A_i^c}$  (where an intersection over an empty set of indices is taken as the whole space). Hence,

$$ET = \sum_{j=2}^{\infty} j P(A_j \cap_{i=2}^{j-1} A_i^c) = \sum_{j=2}^{\infty} \prod_{i=2}^{j-1} (1 - 1/i) \geq C \sum_j \frac{1}{j} = \infty.$$

- Of course, one assumes  $EX_1^2 > 0$ . We did something similar in class, where we proved that  $P(\limsup |S_n|/\sqrt{n} = \infty) = 1$ . If  $X_i$  are symmetric (that is, the law of  $X_1$  is the same as the law of  $-X_1$ ), this immediately implies  $P(\limsup S_n/\sqrt{n} = \infty) \geq 1/2$ , and then the 0-1 law shows that this probability is actually 1. If the law of  $X_1$  is not symmetric, take an independent copy of  $X_i$ , say  $X'_i$ , and apply the conclusion to  $Z_i = X_i - X'_i$ .
- Almost sure convergence of the  $X_i$ 's implies a.s. convergence of the  $S_n/n$ , because for almost every  $\omega$  and every  $\epsilon > 0$  one may find an  $n_0 = n_0(\omega, \epsilon)$  such that for  $n > n_0$ ,  $|X_n| < \epsilon$ . Thus,  $\limsup |S_n|/n < \epsilon$ .  
Convergence in probability of the  $X_i$  does NOT imply the convergence in probability of  $S_n$ . A simple example is to take  $X_i = i^2$  with probability  $1/\log i$  (check that the event  $X_i = i^2$  occurs infinitely many times!).
- This is a simple application of the strong LLN: recall that the  $a_i$  are i.i.d., uniform on  $\{0, 1, 2\}$ . Hence, almost surely, since  $Ea_1 = 1$ ,  $E1_{a_1=0} = 1/3$ ,

$$\limsup n^{-1} \sum_{k=1}^n a_k = \lim n^{-1} \sum_{k=1}^n a_k = 1$$

and

$$\limsup n^{-1} \sum_{k=1}^n 1_{a_k=0} = \lim n^{-1} \sum_{k=1}^n 1_{a_k=0} = 1/3.$$